

A Fast and Efficient Parameter Estimation Algorithm for Generalized Output Error Models

Jie Jia^{1,2}, Hua Huang², Yong Yang^{3,*}, Ke Lv¹, Feng Ding⁴ and Shuying Huang⁵

¹ School of Computing and Communication Engineering, University of Chinese Academy of Sciences, Beijing 100049, China

² Institute of Aerospace Information and Security Technology, Nanchang Hangkong University, Nanchang 330063, China

³ School of Information Technology, Jiangxi University of Finance and Economics, Nanchang 330032, China

⁴ School of Internet of Things Engineering, Jiangnan University, Wuxi 214122, China

⁵ School of Software and Communication Engineering, Jiangxi University of Finance and Economics, Nanchang 330032, China

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Abstract: One kind of the colored noise interference systems is generalized output error model (OEARMA). This paper presents a two-stage recursive least squares algorithm for OEARMA. Aiming at the OEARMA, this paper puts forward a two-stage recursive least squares algorithm. The basic idea of the algorithm is to combine the auxiliary model identification idea and the decomposition technique to decompose a system into two subsystems. Each subsystem contains a parameter vector. With auxiliary model-based recursive extended least squares theory, an unknown intermediate variable output instead of the auxiliary model identification model vector, instead of unmeasurable noise terms in the information vector with the estimated residuals, which can use recursive identification idea to estimate all the parameters of the system, the algorithm has a high computational efficiency. The example of simulation states the effectiveness of the proposed algorithm.

Keywords: Stochastic system, Least squares, Two-stage algorithm, Auxiliary model

1 Introduction

System identification of the colored noise has always been the domestic and foreign scholars are concerned about areas of research [1,2]. Stochastic systems with colored noise, the conventional least squares parameter estimation is biased [2]. Many scholars have done a lot of work, but also made a lot of effective methods for least squares algorithm identification system there is a deviation of colored noise, For example, Bias Compensation Least Squares algorithm (BCLS) [3], Recursive Generalized Least Squares (RGLS) algorithm [4], Recursive Generalized Extended Least Squares (RGELS) algorithm [4], two-stage recursive least squares parameter estimation algorithm [5], and in 1991 the thinking of the auxiliary model identification proposed by Ding Feng [6], the hierarchical identification principle [7], multi-innovation identification theory [8] and parameter estimation error bounds theory [9,10,11,12] and so on.

These methods can not only give the system model parameter estimation, and the latter two can produce noise model parameter estimation.

However, theoretical analysis shows that the unbiased estimate of bias compensation least squares algorithm for OEARMA model is difficult to do, and asked to enter is smooth (Stationary), ergodic (Ergodic). Ding Feng, improved bias compensation least squares algorithm to overcome these shortcomings [13], [14] using the filter to filter the input data, is bound to increase the amount of calculation. RGLS algorithm [15] in the process of the output signal to noise ratio is relatively large or the model parameters for a long time, this white processing of the data reliability will drop. May appear multiple local convergence point recognition accuracy is also low, so that the final identification result is biased. RGELS algorithm convergence as well as the convergence of the theory under what conditions proved challenging subject gives only an approximate analysis of the literature. Feng Ding et al. [16] proposed an iterative identification method identify CARAR model to obtain a satisfactory

* Corresponding author e-mail: yangyong5080@126.com

accuracy, but because of the complexity of practical problems, not online identification, model selection is more difficult to whiten at the same time the algorithm is also more complex, under what conditions convergence is also not a good solution. However, theoretical analysis shows that the RELS convergence of the algorithm requirements for noise model is strictly positive real transfer function [17, 18, 19].

The recursive least squares algorithm to solve the ARX model identification problem [20], the recursive extended least squares algorithm to solve the problem of identification of ARMAX model [20], the auxiliary model identification method [21, 22, 23, 24, 25] and bias compensation method to solve the identification problem of the output error model. Auxiliary variable least squares algorithm can be used to identify the system, but can not be given parameter estimation of the noise model. For output error model, in addition to the above-mentioned method, using the rational fraction equivalence method, and further using of relevant technology, proposed parameter estimation algorithm according to the finite impulse response model order incremental. Using rational fraction equivalence method, studied multi-input single-output system identification problem; using rational fraction equivalent ways to simplify the colored noise stochastic system, use the approximate simplified model can be extended least squares algorithm to estimate the parameters, and then determine the parameters of the original system.

The above-mentioned method only solves the special OEARMA model identification when a polynomial or two polynomials is 1. For the general form of stochastic systems, this paper proposes a class of generalized output error model two-stage recursive extended least squares parameter estimation algorithm. The basic idea of this algorithm is a combination of the auxiliary model identification of ideas and the decomposition technique, the system is decomposed into two subsystems, each subsystem contains a parameter vector.

With auxiliary model and the recursive extended least squares theory, using the auxiliary model's output instead of unknown intermediate variable of the identification model vector, and using the estimated residuals instead of unpredictable noise of the information vector, thus we can use recursive identification ideas estimated all the parameters of the system.

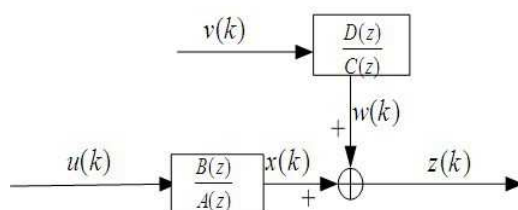


Fig. 1 The Random system block diagram

2 System description and identification model

Defining that “A:=X” or “X:=A” signifies “A is equivalent to X”; I signifies a unit matrix of appropriate sizes ($n \times n$); the superscript T signifies the matrix/vector transpose; I_n signifies an n -dimensional column vector whose elements are 1.

Consider the output error system, described in Fig. 1,

$$z(k) = \frac{B(z)}{A(z)}u(k) + \frac{D(z)}{C(z)}v(k) \quad (1)$$

Among them, $\{u(k)\}$ is a system input at time K , $\{z(k)\}$ is a system output at time K , $\{v(k)\}$ is an uncorrelated random white noise with zero mean and its variance is σ^2 , $A(z)$, $B(z)$, $C(z)$ and $D(z)$ are polynomials in the unit backward shift functor z^{-1} [i.e., $z^{-1}y(k) = y(k-1)$], moreover,

$$A(z) := 1 + a_1z^{-1} + a_2z^{-2} + \dots + a_{n_a}z^{-n_a}$$

$$B(z) := b_1z^{-1} + b_2z^{-2} + \dots + b_{n_b}z^{-n_b}$$

$$C(z) := 1 + c_1z^{-1} + c_2z^{-2} + \dots + c_{n_c}z^{-n_c}$$

$$D(z) := 1 + d_1z^{-1} + d_2z^{-2} + \dots + d_{n_d}z^{-n_d}$$

We may assume that $k \leq 0$, $u(k) = 0$, $z(k) = 0$, $v(k) = 0$, and the order known. The goal of this article is based on the separation of two-stage identification algorithm, the original recognition system into two sub-problems of smaller order. Define parameter vectors,

$$\theta := \begin{bmatrix} \theta_s \\ \theta_n \end{bmatrix} \in R^n, n = n_a + n_b + n_c + n_d$$

$$\theta_s := [a_1, a_2, \dots, a_{n_a}, b_1, b_2, \dots, b_{n_b}]^T \in R^{n_a+n_b}$$

$$\theta_n := [c_1, c_2, \dots, c_{n_c}, d_1, d_2, \dots, d_{n_d}]^T \in R^{n_c+n_d}$$

Define the information vectors,

$$\varphi(k) := \begin{bmatrix} \varphi_s(k) \\ \varphi_n(k) \end{bmatrix} \in R^n, n = n_a + n_b + n_c + n_d$$

$$\varphi_s(k) := [-x(k-1), -x(k-2), \dots, -x(k-n_a),$$

$$u(k-1), u(k-2), \dots, u(k-n_b)] \in R^{(n_a+n_b)}$$

$$\varphi_n(k) := [-w(k-1), -w(k-2), \dots, -w(k-n_c),$$

$$v(k-1), v(k-2), \dots, v(k-n_d)] \in R^{(n_c+n_d)}$$

Define the intermediate variables $x(k)$, $w(k)$, as follows,

$$x(k) := \frac{B(z)}{A(z)}u(k) \quad (2)$$

or

$$\begin{aligned}
 x(k) &= [1 - A(z)]x(k) + B(z)u(k) \\
 &= (-a_1z^{-1} - a_2z^{-2} - \dots - a_{n_a}z^{-n_a})x(k) \\
 &\quad + (b_1z^{-1} + b_2z^{-2} + \dots + b_{n_b}z^{-n_b})u(k) \\
 &= -a_1x(k-1) - a_2x(k-2) - \dots - a_{n_a}x(k-n_a) \\
 &\quad + b_1u(k-1) + b_2u(k-2) + \dots + b_{n_b}u(k-n_b) \\
 &= \varphi_s^T(k)\theta_s \\
 \\
 w(k) &:= \frac{D(z)}{C(z)}v(k) \tag{3}
 \end{aligned}$$

or

$$\begin{aligned}
 w(k) &= [1 - C(z)]w(k) + D(z)v(k) \\
 &= (-c_1z^{-1} - c_2z^{-2} - \dots - c_{n_c}z^{-n_c})w(k) \\
 &\quad + (1 + d_1z^{-1} + d_2z^{-2} + \dots + d_{n_d}z^{-n_d})v(k) \\
 &= -c_1w(k-1) - c_2w(k-2) - \dots - c_{n_c}w(k-n_c) \\
 &\quad + v(k) + d_1v(k-1) + d_2v(k-2) + \dots + d_{n_d}v(k-n_d) \\
 &= \varphi_n^T(k)\theta_n + v(k)
 \end{aligned}$$

Using (2) and (3), (1) can be expressed as

$$\begin{aligned}
 z(k) &= x(k) + w(k) \\
 &= \varphi_s^T(k)\theta_s + \varphi_n^T(k)\theta_n + v(k) \\
 &= \varphi^T(k)\theta + v(k) \tag{4}
 \end{aligned}$$

3 Two-stage recursive least squares algorithm

The basic idea of the algorithm is to transform the system into two subsystems, the parameter vector and the information vector were also transformed into two sub-parameter vectors and two sub-information vectors. Then the auxiliary model identification idea is used to identify the parameters of each subsystem. Define two intermediate variables,

$$z_1(k) := z(k) - \varphi_n^T(k)\theta_n \tag{5}$$

$$z_2(k) := z(k) - \varphi_s^T(k)\theta_s \tag{6}$$

The system in (4) can be transformed into the two virtual identification subsystems, as follows,

$$z_1(k) = \varphi_s^T(k)\theta_s + v(k)$$

$$z_2(k) = \varphi_n^T(k)\theta_n + v(k)$$

These two subsystems contain the parameter vectors θ_s and θ_n , separately. Define two criterion functions, as follows,

$$J_1(\theta_s) := \sum_{j=1}^k [z_1(j) - \varphi_s^T(j)\theta_s]^2$$

$$J_2(\theta_n) := \sum_{j=1}^k [z_2(j) - \varphi_n^T(j)\theta_n]^2$$

Make the partial derivatives of $J_1(\theta_s)$ and $J_2(\theta_n)$ for θ_s and θ_n be zero, separately,

$$\frac{\partial J_1(\theta_s)}{\partial \theta_s} = -2\varphi_s(j) \sum_{j=1}^k [z_1(j) - \varphi_s^T(j)\theta_s] = 0$$

$$\frac{\partial J_2(\theta_n)}{\partial \theta_n} = -2\varphi_n(j) \sum_{j=1}^k [z_2(j) - \varphi_n^T(j)\theta_n] = 0$$

Make $\hat{\theta}(k) := \begin{bmatrix} \hat{\theta}_s(k) \\ \hat{\theta}_n(k) \end{bmatrix} \in R^n$ be the estimate of $\theta := \begin{bmatrix} \theta_s \\ \theta_n \end{bmatrix} \in R^n$ at time k . Then minimizing the criterion functions, so we can get the recursive least squares (RLS) algorithm,

$$\hat{\theta}_s(k) = \hat{\theta}_s(k-1) + K_s(k)[z_1(k) - \varphi_s^T(k)\hat{\theta}_s(k-1)] \tag{7}$$

$$K_s(k) = P_s(k-1)\varphi_s(k)[1 + \varphi_s^T(k)P_s(k-1)\varphi_s(k)]^{-1} \tag{8}$$

$$P_s(k) = [I - K_s(k)\varphi_s^T(k)]P_s(k-1), P_s(0) = P_0I \tag{9}$$

$$\hat{\theta}_n(k) = \hat{\theta}_n(k-1) + K_n(k)[z_2(k) - \varphi_n^T(k)\hat{\theta}_n(k-1)] \tag{10}$$

$$K_n(k) = P_n(k-1)\varphi_n(k)[1 + \varphi_n^T(k)P_n(k-1)\varphi_n(k)]^{-1} \tag{11}$$

$$P_n(k) = [I - K_n(k)\varphi_n^T(k)]P_n(k-1), P_n(0) = P_0I \tag{12}$$

Substituting (5) and (6) into (7) and (10), separately, then

$$\begin{aligned}
 \hat{\theta}_s(k) &= \hat{\theta}_s(k-1) + K_s(k)[z(k) - \varphi_n^T(k)\theta_n \\
 &\quad - \varphi_s^T(k)\hat{\theta}_s(k-1)] \tag{13}
 \end{aligned}$$

$$\begin{aligned}
 \hat{\theta}_n(k) &= \hat{\theta}_n(k-1) + K_n(k)[z(k) - \varphi_s^T(k)\theta_s \\
 &\quad - \varphi_n^T(k)\hat{\theta}_n(k-1)] \tag{14}
 \end{aligned}$$

Using the estimates $\hat{\theta}_n(k-1)$ and $\hat{\theta}_s(k-1)$ to replace the unknown parameter vectors at the right-hand sides of (13) and (14), separately, then we can get

$$\begin{aligned}
 \hat{\theta}_s(k) &= \hat{\theta}_s(k-1) + K_s(k)[z(k) - \varphi_n^T(k)\hat{\theta}_n(k-1) \\
 &\quad - \varphi_s^T(k)\hat{\theta}_s(k-1)] \tag{15}
 \end{aligned}$$

$$\begin{aligned}
 \hat{\theta}_n(k) &= \hat{\theta}_n(k-1) + K_n(k)[z(k) - \varphi_s^T(k)\hat{\theta}_s(k-1) \\
 &\quad - \varphi_n^T(k)\hat{\theta}_n(k-1)] \tag{16}
 \end{aligned}$$

Using the estimates $\varphi_s(k)$ and $\varphi_n(k)$ to replace the unknown information vectors at the right-hand sides of (8), (11), (15) and (16). Finally, we can get

$$\begin{aligned}
 \hat{\varphi}_s(k) &:= [-\hat{x}(k-1), -\hat{x}(k-2), \dots, -\hat{x}(k-n_a), \\
 &\quad \hat{u}(k-1), \hat{u}(k-2), \dots, \hat{u}(k-n_b)] \in R^{n_a+n_b} \tag{17}
 \end{aligned}$$

$$\begin{aligned}
 \hat{\varphi}_n(k) &:= [-\hat{w}(k-1), -\hat{w}(k-2), \dots, -\hat{w}(k-n_c), \\
 &\quad \hat{v}(k-1), \hat{v}(k-2), \dots, \hat{v}(k-n_d)] \in R^{n_c+n_d} \tag{18}
 \end{aligned}$$

Define,

$$\hat{\varphi}(k) := \begin{bmatrix} \hat{\varphi}_s(k) \\ \hat{\varphi}_n(k) \end{bmatrix} \in R^n$$

Replacing $\varphi_s(k), \varphi_n(k), \theta_s(k), \theta_n(k)$ in (2), (3) and (4) with $\hat{\varphi}_s(k), \hat{\varphi}_n(k), \hat{\theta}_s(k), \hat{\theta}_n(k)$, giving,

$$\begin{aligned} \hat{x}(k) &= \hat{\varphi}_s^T(k)\hat{\theta}_s, \\ \hat{w}(k) &= z(k) - \hat{\varphi}_s^T(k)\hat{\theta}_s, \\ \hat{v}(k) &= z(k) - \hat{\varphi}^T(k)\hat{\theta} \end{aligned} \tag{19}$$

We can obtain the two-stage recursive least squares identification algorithm (TS-RLS) for estimating the parameter vectors θ_n and θ_s of the OE models, as follows,

$$\begin{aligned} \hat{\theta}_s(k) &= \hat{\theta}_s(k-1) + K_s(k)[z(k) - \hat{\varphi}_n^T(k)\hat{\theta}_n(k-1) \\ &\quad - \hat{\varphi}_s^T(k)\hat{\theta}_s(k-1)] \tag{20} \\ K_s(k) &= P_s(k-1)\hat{\varphi}_s(k)[1 + \hat{\varphi}_s^T(k)P_s(k-1)\hat{\varphi}_s(k)]^{-1} \tag{21} \\ P_s(k) &= [I - K_s(k)\hat{\varphi}_s^T(k)]P_s(k-1), P_s(0) = P_0I \tag{22} \\ \hat{\varphi}_s(k) &= [-\hat{x}(k-1), -\hat{x}(k-2), \dots, -\hat{x}(k-n_a), \\ &\quad u(k-1), u(k-2), \dots, u(k-n_b)] \tag{23} \\ \hat{\theta}_n(k) &= \hat{\theta}_n(k-1) + K_n(k)[z(k) - \hat{\varphi}_s^T(k)\hat{\theta}_s(k) \\ &\quad - \hat{\varphi}_n^T(k)\hat{\theta}_n(k)] \tag{24} \\ K_n(k) &= P_n(k-1)\hat{\varphi}_n(k)[1 + \hat{\varphi}_n^T(k)P_n(k-1)\hat{\varphi}_n(k)]^{-1} \tag{25} \\ P_n(k) &= [I - K_n(k)\hat{\varphi}_n^T(k)]P_n(k-1), P_n(0) = P_0I \tag{26} \\ \hat{\varphi}_n(k) &= [-\hat{w}(k-1), -\hat{w}(k-2), \dots, -\hat{w}(k-n_c), \\ &\quad \hat{v}(k-1), \hat{v}(k-2), \dots, \hat{v}(k-n_d)] \tag{27} \\ \hat{x}(k) &= \hat{\varphi}_s^T(k)\hat{\theta}_s, \hat{w}(k) = z(k) - \hat{\varphi}_s^T(k)\hat{\theta}_s, \\ \hat{v}(k) &= z(k) - \hat{\varphi}^T(k)\hat{\theta} \end{aligned} \tag{28}$$

$K_s(k)$ and $K_n(k)$ are two gain vectors, and $P_s(k)$ and $P_n(k)$ are two covariance matrices.

The steps of computing $\hat{\theta}_s(k), \hat{\theta}_n(k)$ in the RLS algorithm in (20)–(28) are listed in the following:

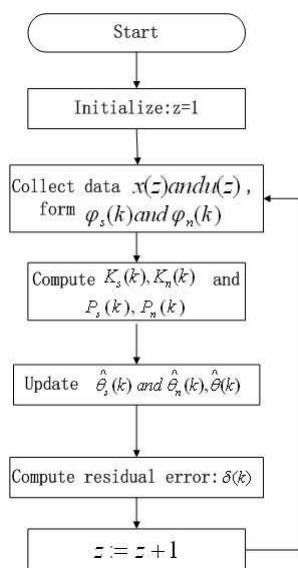


Fig. 2 The flowchart of computing the OEARMA parameter estimation $\hat{\theta}_s(k), \hat{\theta}_n(k)$

4 The recursive extended least squares algorithm-RELS

To compare with the proposed algorithm, the auxiliary model based recursive least square algorithm is introduced in this section. Recursive Extended Least Squares algorithm is an identification method, which is used to deal with colored noise of the CARMA model by increasing the dimension of the parameter vector and dope-vector. That is to say, noise regression item is added in the information vector, while noise model parameters are mixed in the parameter vector.

$$\begin{aligned} \theta &:= [a_1, a_2, \dots, a_{n_a}, b_1, b_2, \dots, b_{n_b}, c_1, c_2, \dots, c_{n_c}, \\ &\quad d_1, d_2, \dots, d_{n_d}] \in R^{n_a+n_b+n_c+n_d} \\ \varphi(k) &:= [-z(k-1), -z(k-2), \dots, -z(k-n_a), u(k-1), \\ &\quad u(k-2), \dots, u(k-n_b), -w(k-1), -w(k-2), \dots, \\ &\quad -w(k-n_c), v(k-1), v(k-2), \dots, v(k-n_d)] \\ J(\theta) &:= \sum_{j=1}^k [z(k) - \varphi^T(k)\theta]^2 \\ \hat{\theta}(k) &= \hat{\theta}(k-1) + P(k)\hat{\varphi}(k)[z(k) - \hat{\varphi}^T(k)\hat{\theta}(k-1)] \\ P(k) &= P(k-1) - \frac{P(k-1)\hat{\varphi}(k)\hat{\varphi}^T(k)P(k-1)}{1 + \hat{\varphi}^T(k)P(k-1)\hat{\varphi}(k)}, P(0) = p_0I \\ \hat{v}(k) &= z(k) - \hat{\varphi}^T(k)\hat{\theta}(k-1) \\ \hat{\varphi}(k) &= [-z(k-1), -z(k-2), \dots, -z(k-n_a), u(k-1), \\ &\quad u(k-2), \dots, u(k-n_b), -\hat{w}(k-1), -\hat{w}(k-2), \dots, \\ &\quad -\hat{w}(k-n_c), \hat{v}(k-1), \hat{v}(k-2), \dots, \hat{v}(k-n_d)] \\ \hat{\theta} &= [\hat{a}_1, \hat{a}_2, \dots, \hat{a}_{n_a}, \hat{b}_1, \hat{b}_2, \dots, \hat{b}_{n_b}, \hat{c}_1, \hat{c}_2, \dots, \hat{c}_{n_c}, \\ &\quad \hat{d}_1, \hat{d}_2, \dots, \hat{d}_{n_d}] \end{aligned}$$

5 Example

Consider the following OEARMA system,

$$\begin{aligned} Z(k) &= \frac{B(k)}{A(k)}u(k) + \frac{D(k)}{C(k)}v(k) \\ w(k) &= \frac{D(k)}{C(k)}v(k) \\ A(z) &= 1 + a_1z^{-1} + a_2z^{-2} = 1 + 1.60z^{-1} + 0.8z^{-2} \\ B(z) &= b_1z^{-1} + b_2z^{-2} = 0.412z^{-1} + 0.309z^{-2} \\ C(z) &= 1 + c_1z^{-1} = 1 + 0.8z^{-1} \\ D(z) &= 1 + d_1z^{-1} = 1 - 0.64z^{-1} \\ \theta^T &= [1.60, 0.8, 0.412, 0.309, 0.8, -0.64] \end{aligned}$$

In simulation, the system input $\{u(k)\}$ is an uncorrelated random signal sequence with zero mean and unit variance, and $\{v(k)\}$ as a white noise sequence with zero mean and its variance is σ^2 . Using the RLS algorithm to estimate the parameters of this system, the

parameter estimate and estimation errors $\delta := \frac{\|\hat{\theta} - \theta\|}{\|\theta\|}$ versus k are presented in Tables 1 and 2. The parameter estimation errors of the RLS algorithm with different variances are presented in Fig. 3. The parameter estimation errors for the two algorithms are presented in Fig. 4. From these Tables and Figs, we can find that:

- With the gradual increasing of noise variance decreases, the precision of the parameter estimates gradually improved, referring to Fig.3 to control;
- With the increase of the length of data, the parameter estimation error is smaller and smaller, and in Fig.4, the RLS algorithm can visually see superior RELS algorithm;
- The reference to Table 3 that the amount of calculation of the RLS algorithm respect RELS algorithm to be greatly simplified, wherein $n = n_a + n_b + n_c + n_d$.

Table 1 RLS algorithm parameter estimates and residuals

k	a_1	a_2	b_1	b_2	c_1	d_1	$\delta(\%)$
100	1.20260	0.50691	0.44384	0.10396	0.79078	-0.61422	25.24060
200	1.38256	0.62701	0.43945	0.20664	0.80789	-0.73193	14.65398
500	1.50781	0.73472	0.41297	0.25746	0.82262	-0.63651	5.94194
1000	1.56554	0.77144	0.42204	0.29850	0.79953	-0.61600	2.48625
2000	1.56720	0.77511	0.40969	0.30610	0.79876	-0.63711	1.95130
3000	1.57315	0.77594	0.41251	0.30417	0.79270	-0.67137	2.28639
4000	1.61920	0.81367	0.40720	0.32474	0.79790	-0.66544	1.80940
5000	1.60369	0.79932	0.40738	0.31820	0.79974	-0.66888	1.45374
6000	1.61022	0.80409	0.40868	0.32003	0.80252	-0.65743	1.11775
7000	1.60706	0.80172	0.41220	0.31841	0.80221	-0.66047	1.11907
True values	1.60000	0.80000	0.41200	0.30900	0.80000	-0.64000	0.00000

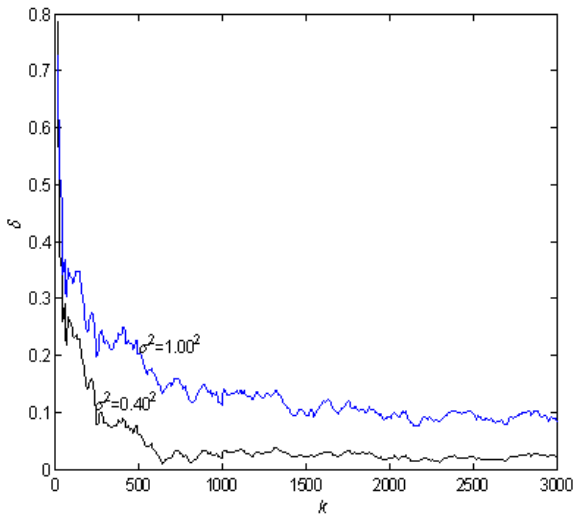


Fig. 3 The RLS parameter estimation errors δ versus k ($\sigma^2 = 0.40^2$ and $\sigma^2 = 1.00^2$)

Table 2 RLS algorithm parameter estimates and residuals

k	a_1	a_2	b_1	b_2	c_1	d_1	$\delta(\%)$
100	1.64648	0.77137	0.42953	0.30980	0.54707	-0.60255	12.33223
200	1.63711	0.78087	0.41288	0.32958	0.68643	-0.68990	6.23565
500	1.59082	0.74141	0.40771	0.28768	0.77035	-0.65435	3.35251
1000	1.58187	0.73913	0.42097	0.30178	0.76132	-0.65371	3.60007
2000	1.56228	0.72573	0.40866	0.30378	0.78236	-0.68412	4.52285
3000	1.55633	0.72314	0.41106	0.29813	0.78266	-0.72008	5.69570
4000	1.54888	0.72009	0.40550	0.29869	0.80517	-0.70123	5.35111
5000	1.54388	0.71793	0.40621	0.29756	0.81598	-0.70100	5.57394
6000	1.54133	0.71753	0.40815	0.29671	0.82358	-0.68668	5.39588
7000	1.53665	0.71462	0.41102	0.29458	0.82989	-0.68782	5.70455
True values	1.60000	0.80000	0.41200	0.30900	0.80000	-0.64000	0.00000

Table 3 Comparison of the computational efficiency of the RLS and RELS algorithms

Computational efficiency \ Algorithms	RLS	RELS
Number of multiplication	$2(n_a + n_b)^2 + 2(n_c + n_d)^2 + 4n$	$2n^2 + 4n$
Number of additions	$2(n_a + n_b)^2 + 2(n_c + n_d)^2 + 2n$	$2n^2 + 2n$
The total number of calculation	$4(n_a + n_b)^2 + 4(n_c + n_d)^2 + 6n$	$4n^2 + 6n$

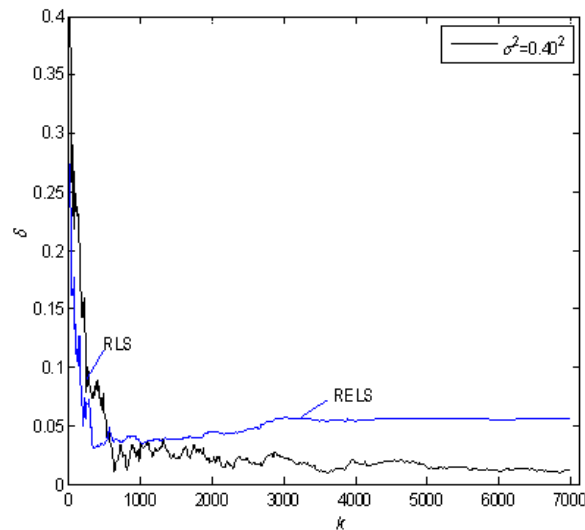


Fig. 4 The parameter estimation errors versus t for the two algorithms ($\sigma^2 = 0.40^2$)

6 Conclusion

A class of generalized output error model of two-stage recursive least squares parameter estimation algorithm has been derived in the paper. with the help of the auxiliary model identification idea and decomposition techniques ,the system is converted to the two-step process in the proposed algorithm. Use the measurable information of the system to build a auxiliary model, of which inputs can replace the unmeasurable variates. By choosing the parameters of the auxiliary model, the inputs of the auxiliary model can approach to these unmeasurable variates, thus obtaining the concordant estimation of the system variates. The derivation of the algorithm is simple, less computation, high accuracy. Theoretical analysis and simulation results verify these conclusions.

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Jie Jia received his Master degree and PH.D degree at Northwestern Polytechnical University. Currently he is a professor of the School of Information Engineering at Nanchang Aeronautical University. He has published more than 30

core papers. His research focuses on aircraft navigation guidance and control, test technology research of aeronautics and astronautics, aviation micro sensor system, the control of nonlinear system and its simulation, information modeling and simulation technology.



Hua Huang was born in Quannan Jiangxi on May 27, 1990 and graduated from School of Information Engineering at Nanchang Aeronautical University in July 2012. Currently she is a postgraduate and her major is Control Engineering for

the School of Information Engineering at Nanchang Aeronautical University. Her research interests include intelligent control and navigation guidance and control.



Yong Yang received his Ph.D. degree in Biomedical Engineering from Xi'an Jiaotong University, China, in 2005. From 2009 to 2010, he was a postdoctoral research fellow at the Chonbuk National University, Republic of Korea. He is

currently a Full Professor with the School of Information Technology, Jiangxi University of Finance and Economics, China. He has published more than 70 research articles in international journals and conferences. His research interests include image processing, medical image analysis, signal processing, and pattern recognition.



Ke Lv was born in Guyuan Ningxia on March 13, 1971, he graduated from Department of Mathematics at Ningxia University in July 1993. He received Master degree and PH.D degree from Department of Mathematics and Department of Computer science at Northwest University in July 1998 and

July 2003, respectively. He as Postdoctoral Fellow in Institute of Automation Chinese Academy of Sciences from July 2003 to April 2005. Currently he is a professor of University of the Chinese Academy of Sciences. He research focuses on curve matching, 3D image reconstruction and computer graphics.



Feng Ding was born in Guangshui Hubei on 1963. He graduated from Hubei Polytechnical University in July 1984. He received Master degree and PH.D degree from Tsinghua university in July 1990 and in July 1994. He as Postdoctoral

research fellow in University of Alberta from July 2003 to October 2005, in The Hong Kong university of science and technology from March 2006 to May 2006 and Ryerson University from January 2009 to October 2009. Currently he is a professor of the Internet of Things Institute in Jiangnan University. He has published more than 200 Academic papers in domestic and international professional journals and magazines. His research focuses on system modeling, system identification, process control, adaptive control.



Shuying Huang received her Ph.D. degree in Computer Application Technology from Ocean University of China, Shandong, China, 2013. She is now a Lecturer with the School of Software and Communication Engineering, Jiangxi University of Finance and Economics, Nanchang, China. She has published

more than 20 academic papers. Her research interests include image and signal processing, and pattern recognition.